

Digital Communications (ELC 623)

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Postgraduate Program

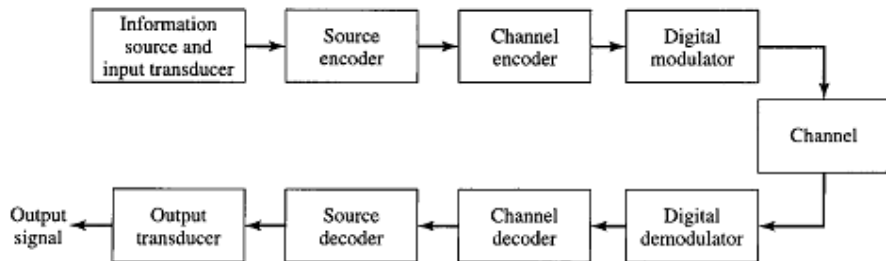
Introduction

Statistical Decision Theory

- 1 Introduction to Communication Systems
- 2 Review on Probability Theory
- 3 Review on Stochastic Processes

Elements of Communication Systems

Transmitter



Receiver

Transmitter

① **Information Source**

Analog - Digital

② **Source Encoder**

Represent the source signal as efficiently as possible (minimize the redundancy)

③ **Channel Encoder**

Increase reliability of received data (add redundancy in a controlled manner to information bits)

④ **Digital Modulator**

Transmit most efficiently over the (physical) transmission channel (map the input bit sequence to a signal waveform which is suitable for the transmission channel)

Receiver

① **Digital Demodulator**

Reconstruct transmitted data symbols (binary or M-ary) from channel-corrupted received signal

② **Channel Decoder**

Exploit redundancy introduced by channel encoder to increase reliability of information bits

③ **Source Decoder**

Reconstruct original information signal from output of channel decoder

Note: In advanced receivers, demodulation and decoding are sometimes performed iteratively to enhance the receiver's performance.

Channel

1 Examples of Physical Channel

- Wireline - Optical Fiber
- Wireless radio frequency (RF) channel - Optical Wireless Channel
- Underwater Acoustic Channel - Storage Channel (CD, disc, etc.)

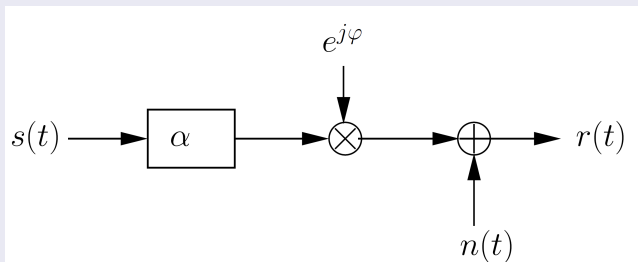
2 Channel Impairments

- Noise (electronic, thermal, ...)
- Nonlinearities, distortions, time-variance, ...
- Interference

For the design of the transmitter and the receiver we need a simple mathematical model of the physical communication channel that captures its most important properties. This model will vary from one application to another.

Additive White Gaussian Noise Channel - With Unknown Phase

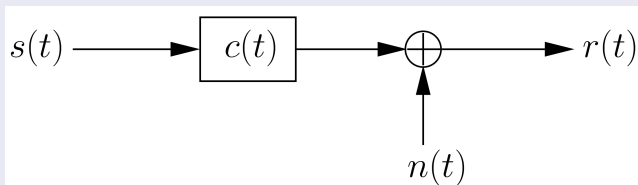
$$r(t) = \alpha e^{j\phi} s(t) + n(t)$$



The transmitted signal experiences an unknown phase shift ϕ , which is often modeled as a random variable, uniformly distributed in the interval $[-\pi, \pi)$. The transmitted signal is also attenuated by a factor of α , and impaired by AWGN.

Linearly Dispersive Channel (Linear Filter Channel)

$$r(t) = c(t) * s(t) + n(t)$$



The transmitted signal is linearly distorted by $c(t)$ and impaired by AWGN.

Types of Channels

- Multiuser channels
- MIMO channels
- Relaying channels
- Fading channels
- ...

Discussion:

What tools do we need for the analysis/design of communication systems?

The Theory of Probability: Random Experiments

An essential tool in the design of digital communication systems.

Probability

- Experiment
- Outcome
- Event
- Probability - Axioms of Probability
- Union - Intersection
- Joint events and joint probabilities
- Conditional probability
- Statistical Independence

Axioms of Probability

Assume events A and B are subsets of the sample space S , i.e. $A \subset S$ and $B \subset S$

- 1 $P(S) = 1$
- 2 $0 < P(A) < 1$
- 3 If $A \cap B = \phi$, then $P(A \cup B) = P(A) + P(B)$

Note: If $A \cap B \neq \phi$, then $P(A \cup B) = P(A) + P(B) - P(A \cap B)$

Joint Events and Joint Probability

Consider two experiments with outcomes $A_i, i = 1, 2, \dots, n$ and $B_j, j = 1, 2, \dots, m$

If both experiments are carried out, then the outcome (A_i, B_j) is assigned the probability $P(A_i, B_j)$ with $0 \leq P(A_i, B_j) \leq 1$

- 1 If the outcome of $B_j, j = 1, 2, \dots, m$ are mutually exclusive, then

$$\sum_{j=1}^m P(A_i, B_j) = P(A_i)$$

- 2 If all the outcome of both $A_i, i = 1, 2, \dots, n$ and $B_j, j = 1, 2, \dots, m$ are mutually exclusive, then

$$\sum_{i=1}^n \sum_{j=1}^m P(A_i, B_j) = 1$$

Conditional Probability

The conditional probability $P(A|B)$ is the probability of event A given that event B has already been observed.

Conditional Probability

$$P(A|B) = \frac{P(A, B)}{P(B)}$$

$$P(B|A) = \frac{P(A, B)}{P(A)}$$

Bayes' Theorem

$$P(A|B) = \frac{P(B|A)P(A)}{P(B)}$$

$$P(B|A) = \frac{P(A|B)P(B)}{P(A)}$$

Statistical Independence

If observing B does not change the probability of observing A, i.e., $P(A|B) = P(A)$, then A and B are statistically independent. In this case:

$$P(A, B) = P(A|B)P(B) = P(A)P(B)$$

Statistically Independent Events

Two events A and B are statistically independent if and only if

$$P(A, B) = P(A)P(B)$$

The Theory of Probability: Random Experiments

Probability

- Experiment
- Outcome
- Event
- Probability - Axioms of Probability
- Union - Intersection
- Joint events and joint probabilities
- Conditional probability
- Statistical Independence

Let's play

Where is the Prize?

Random Variables

- Probability distributions
- Probability densities
- Joint probability distributions
- Conditional probability distributions
- Statistically independent random variables
- Statistical averages

Transformation of Random Variables

$$Y = g(X)$$

Random Variables: Cumulative Distribution Function

The CDF $F(x)$ denotes the probability that a random variable (RV) X is smaller than or equal to a specific value x , i.e.

$$F(x) = P(X \leq x)$$

Properties of CDF

$$\begin{aligned}0 &\leq F(x) \leq 1 \\ \lim_{x \rightarrow -\infty} F(x) &= 0 \\ \lim_{x \rightarrow \infty} F(x) &= 1 \\ \frac{d}{dx} F(x) &\geq 0\end{aligned}$$

Random Variables: Probability Density Function

The PDF of a RV X is defined as:

$$p(x) = \frac{dF(x)}{dx}, \quad -\infty \leq x \leq \infty$$

Properties of PDF

$$p(x) \geq 0$$

$$F(x) = \int_{-\infty}^x p(u) du$$

$$\int_{-\infty}^{\infty} p(u) du = 1$$

Discrete Random Variables

For discrete random variables, where $X \in \{x_1, x_2, \dots, x_n\}$,

$$p(x) = \sum_{i=1}^n P(X = x_i) \delta(x - x_i)$$

Note: The probability that $x_1 \leq X \leq x_2$ is given as

$$P(x_1 \leq X \leq x_2) = \int_{x_1}^{x_2} p(u) du = F(x_2) - F(x_1)$$

Joint CDF and PDF

Given two RVs, X and Y ,

$$\begin{aligned}F_{XY}(x, y) &= P(X \leq x, Y \leq y) \\ &= \int_{-\infty}^x \int_{-\infty}^y p_{XY}(u, v) du dv \\ p_{XY}(x, y) &= \frac{\partial}{\partial x} \frac{\partial}{\partial y} F_{XY}(x, y)\end{aligned}$$

Marginal Densities

$$\begin{aligned}p_X(x) &= \int_{-\infty}^{\infty} p_{XY}(x, y) dy \\ p_Y(y) &= \int_{-\infty}^{\infty} p_{XY}(x, y) dx\end{aligned}$$

Properties of Joint CDF and PDF

$$F_{XY}(-\infty, y) = F_{XY}(x, -\infty) = F_{XY}(-\infty, -\infty) = 0$$
$$F_{XY}(\infty, \infty) = 1$$

Conditional PDF and CDF

$$p_{X|Y}(x|y) = \frac{p_{XY}(x, y)}{p_Y(y)}$$
$$F_{X|Y}(x|y) = \int_{-\infty}^x p_{X|Y}(u|y) du$$

Statistical Independence

X and Y are statistically independent iff

$$p_{XY}(x, y) = p_X(x)p_Y(y)$$

Complex Random Variables

For a complex RV $Z = X + jY$,

CDF:

$$F_Z(z) = P(X \leq x, Y \leq y) = F_{XY}(x, y)$$

PDF

$$p_Z(z) = p_{XY}(x, y)$$

Functions of Random Variables

Problem Statement

Given a RV X with known PDF, $p_X(x)$, and CDF, $F_X(x)$, what are the PDF and the CDF of another RV Y , where $Y = g(X)$

Transformation of RVs

$$p_Y(y) = \frac{p_X(f(y))}{J(x)}, \quad \text{where} \quad J(x) = \left. \frac{dy}{dx} \right|_{x=f(y)}$$

Examples

- 1 $Y = aX + b$
- 2 $Y = aX^2 + b$
- 3 $Y = X_1 + X_2$

Sum of Two Random Variables

Problem

Given two random variables X_1 and X_2 , with joint probability $p_{X_1, X_2}(x_1, x_2)$, what the PDF of $Y = X_1 + X_2$?

Solution

Since $x_1 = y - x_2$, then the PDF of Y and X_2 is obtained as

$$p_{Y, X_2}(y, x_2) = p_{X_1, X_2}(x_1, x_2)|_{x_1=y-x_2} = p_{X_1, X_2}(y - x_2, x_2)$$

Then, using marginal densities, the PDF of Y can be obtained as

$$\begin{aligned} p_Y(y) &= \int_{-\infty}^{\infty} p_{X_1, X_2}(y - x_2, x_2) dx_2 \\ &= \int_{-\infty}^{\infty} p_{X_1, X_2}(x_1, y - x_1) dx_1 \end{aligned}$$

Sum of Two Random Variables

Special Case: Sum of Two SI RVs

If X_1 and X_2 are statistically independent, then

$$p_{X_1, X_2}(x_1, x_2) = p_{X_1}(x_1)p_{X_2}(x_2)$$

Then,

$$\begin{aligned} p_Y(y) &= \int_{-\infty}^{\infty} p_{X_1}(y - x_2)p_{X_2}(x_2)dx_2 \\ &= p_{X_1}(x_1) * p_{X_2}(x_2) \end{aligned}$$

The PDF of Y is the **convolution** of the PDFs of X_1 and X_2 .

Statistical Averages of RVs

General Case: Statistical Averaging

$$\mu = \mathcal{E}\{g(X)\} = \int_{-\infty}^{\infty} g(X)p_X(x)dx$$

Mean

$$\mathcal{E}\{X\} = \int_{-\infty}^{\infty} x p_X(x)dx$$

nth Moment

$$\mathcal{E}\{X^n\} = \int_{-\infty}^{\infty} x^n p_X(x)dx$$

Statistical Averages of RVs

nth Central Moment

$$\mathcal{E}\{(X - \mu)^n\} = \int_{-\infty}^{\infty} (x - \mu)^n p_X(x) dx$$

Variance = 2th Central Moment

$$\begin{aligned}\sigma^2 &= \mathcal{E}\{(X - \mu)^2\} = \int_{-\infty}^{\infty} (x - \mu)^2 p_X(x) dx \\ &= \mathcal{E}\{X^2\} - (\mathcal{E}\{X\})^2\end{aligned}$$

Characteristic Function

$$\psi(jt) = \mathcal{E}\{e^{jtx}\} = \int_{-\infty}^{\infty} e^{jtx} p_X(x) dx$$

Characteristic Function: Properties

$$\psi(jt) = \mathcal{F}_{p_X(x)}(-jt)$$

$$p_X(x) = \frac{1}{2\pi} \int_{-\infty}^{\infty} \psi(jt) e^{-jtx} dt$$

$$\mathcal{E}\{X^n\} = (-j)^n \frac{d^n}{dt^n} \psi(jt) \Big|_{t=0}$$

Sum of Two SI RVs: $Y = X_1 + X_2$

$$\begin{aligned}\psi_Y(jt) &= \mathcal{E}\{e^{jtY}\} \\ &= \mathcal{E}\{e^{jt(X_1+X_2)}\} \\ &= \mathcal{E}\{e^{jt(X_1)} e^{jt(X_2)}\} \\ &= \mathcal{E}\{e^{jt(X_1)}\} \mathcal{E}\{e^{jt(X_2)}\} \\ &= \psi_{X_1}(jt) \psi_{X_2}(jt)\end{aligned}$$

Random Variables

- Probability distributions
- Probability densities
- Joint probability distributions
- Conditional probability distributions
- Statistically independent random variables
- Statistical averages

Transformation of Random Variables

$$Y = g(X)$$

Useful Probability Distributions

- Uniform distribution
- **Gaussian (normal) distribution**
- Rayleigh distribution
- Nakagami- m distribution
- Rician distribution
- Chi-square distribution

Bounding

- Chernoff bound
- Central Limit Theorem

Gaussian Distribution

The Gaussian distribution is an important probability distribution in practice because many physical phenomena can be described by a Gaussian distribution, e.g. AWGN

One-Dimensional Gaussian RV

$$p(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$
$$F(x) = \frac{1}{2} + \frac{1}{2} \operatorname{erf} \left(\frac{x-\mu}{\sqrt{2}\sigma} \right)$$

Note: Useful definitions

$$\operatorname{erf}(x) = \frac{2}{\sqrt{\pi}} \int_0^x e^{-t^2} dt$$
$$Q(x) = \frac{1}{\sqrt{2\pi}} \int_x^{\infty} e^{-\frac{t^2}{2}} dt$$

Gaussian Distribution: Statistical Averages

One-Dimensional Gaussian RV

$$\begin{aligned}\psi(jt) &= e^{jt\mu - t^2\sigma^2/2} \\ \mu_k = \mathcal{E}\{(x - \mu)^k\} &= \begin{cases} 1.3.5 \dots (k-1)\sigma^k, & \text{even } k \\ 0, & \text{odd } k \end{cases} \\ \mathcal{E}\{x^k\} &= \sum_{i=0}^k \binom{k}{i} \mu^i \mu_{k-i}\end{aligned}$$

Sum of n SI Gaussian RVs

$$\begin{aligned}Y &= \sum_{i=1}^n X_i \\ \psi_Y(jt) &= \dots\end{aligned}$$

Useful Probability Distributions

- Uniform distribution
- Rayleigh distribution
- Nakagami- m distribution
- Rician distribution
- Chi-square distribution

Refer to [Proakis, Section 2.3]

Tail Probability

The tail probability (area under the tail of PDF) arises often to determine the error probability of digital communication systems, and is given by

$$\begin{aligned} P(x \geq \delta) &= \int_{\delta}^{\infty} p(x) dx \\ &= \int_{-\infty}^{\infty} U(x - \delta) p(x) dx = \mathcal{E}\{U(x - \delta)\} \end{aligned}$$

Note that, for any $\alpha \geq 0$,

$$U(x - \delta) \leq e^{\alpha(x - \delta)}$$

Chernoff Bound

The tail probability (area under the tail of PDF) arises often to determine the error probability of digital communication systems, and is given by

$$\begin{aligned}P(x \geq \delta) &= \mathcal{E}\{U(x - \delta)\} \\ &\leq \mathcal{E}\{e^{\alpha(x-\delta)}\} \\ &= e^{-\alpha\delta} \mathcal{E}\{e^{\alpha x}\}\end{aligned}$$

In order to obtain the tightest Chernoff bound, α should be optimized such that

$$\frac{d}{d\alpha} e^{-\alpha\delta} \mathcal{E}\{e^{\alpha x}\} = 0$$

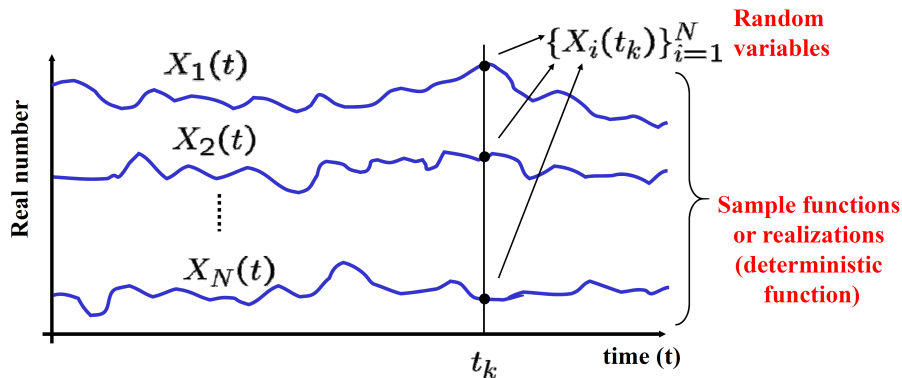
Refer to [Proakis, Section 2.5]

Stochastic processes arise whenever a random phenomenon is a function of time

Stochastic Processes

- Stationary processes - Wide sense stationary processes
- Statistical averages
- Power Spectral Density
- Ergodic processes
- Cyclo-stationary processes

Refer to [Proakis, Section 2.7]



Random Process - Sample Function - Ensemble

Strict Sense Stationary Stochastic Process

RP \rightarrow RV

Considering specific time instants $t_1 > t_2 > \dots > t_n$ with the arbitrary positive integer index n , the random variables $X_{t_i} = X(t_i)$, $i = 1, 2, \dots, n$, are fully characterized by their joint PDF $p(x_{t_1}, x_{t_2}, \dots, x_{t_n})$

Definition

If X_{t_i} and $X_{t_i+\tau}$ have the same statistical properties, $X(t)$ is stationary in the strict sense.

$$p(x_{t_1}, x_{t_2}, \dots, x_{t_n}) = p(x_{t_1+\tau}, x_{t_2+\tau}, \dots, x_{t_n+\tau})$$

Statistical Averages - Ensemble Averages

First-Order Moment = Mean

$$m(t_i) = \mathcal{E}\{X_{t_i}\} = \int_{-\infty}^{\infty} x_{t_i} p(x_{t_i}) dx_{t_i}$$

For SSS RP, $m(t_i) = m$

Second-Order Moment = Autocorrelation Function

$$\phi(t_1, t_2) = \mathcal{E}\{X_{t_1} X_{t_2}\} = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} x_{t_1} x_{t_2} p(x_{t_1}, x_{t_2}) dx_{t_1} dx_{t_2}$$

For SSS RP, $\phi(t_1, t_2) = \phi(t_1 - t_2) = \phi(\tau) \Rightarrow$ **Average Power** = $\phi(0)$

Central Second-Order Moment = Covariance Function

$$\mu(t_1, t_2) = \mathcal{E}\{(X_{t_1} - m(t_1))(X_{t_2} - m(t_2))\} = \phi(t_1, t_2) - m(t_1)m(t_2)$$

For SSS RP, $\mu(t_1, t_2) = \mu(\tau) = \phi(\tau) - m^2 \Rightarrow$ **Variance** = $\mu(0)$

Definition

If the first and second order moments of a stochastic process are invariant to any time shift τ , the process is referred to as wide sense stationary process.

Wide sense stationary processes are not necessarily strict sense stationary.

Definition

A process $X(t)$ is ergodic if its statistical averages can be calculated as timeaverages of sample functions.

Only wide sense stationary processes can be ergodic.

$$m = \lim_{T \rightarrow \infty} \frac{1}{2T} \int_{-T}^T x(t) dt$$
$$\phi(\tau) = \lim_{T \rightarrow \infty} \frac{1}{2T} \int_{-T}^T x(t)x(t + \tau) dt$$

where $x(t)$ is any of the sample functions.

Jointly Stochastic Processes

Joint Stationarity

$X(t)$ and $Y(t)$ are jointly stationary if their joint PDF is invariant to time shifts, τ .

Cross Correlation Function

$$\phi_{XY}(t_1, t_2) = \mathcal{E}\{X_{t_1} Y_{t_2}\} = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} x_{t_1} y_{t_2} p(x_{t_1}, y_{t_2}) dx_{t_1} dy_{t_2}$$

If $X(t)$ and $Y(t)$ are jointly and individually SSS, $\phi_{XY}(t_1, t_2) = \phi_{XY}(\tau)$

Cross Covariance Function

$$\begin{aligned} \mu_{XY}(t_1, t_2) &= \mathcal{E}\{(X_{t_1} - m_X(t_1))(Y_{t_2} - m_Y(t_2))\} \\ &= \phi_{XY}(t_1, t_2) - m_X(t_1)m_Y(t_2) \end{aligned}$$

If $X(t)$ and $Y(t)$ are jointly and individually SSS, $\mu_{XY}(t_1, t_2) = \mu_{XY}(\tau)$

Statistical Independence

$X(t)$ and $Y(t)$ are SI iff for every n and m

$$p(x_{t_1}, x_{t_2}, \dots, x_{t_n}, y_{t_1}, y_{t_2}, \dots, y_{t_m}) = p(x_{t_1}, x_{t_2}, \dots, x_{t_n})p(y_{t_1}, y_{t_2}, \dots, y_{t_m})$$

Uncorrelation

$X(t)$ and $Y(t)$ are uncorrelated iff

$$\mu_{XY}(t_1, t_2) = 0$$

Power Spectral Density

Definition

Power spectral density is defined for stationary RP as the Fourier Transform of the ACF

$$\Phi(f) = \mathcal{F}\{\phi(\tau)\} = \int_{-\infty}^{\infty} \phi(\tau) e^{-j2\pi f\tau} d\tau$$

$$\phi(\tau) = \mathcal{F}^{-1}\{\Phi(f)\} = \int_{-\infty}^{\infty} \Phi(f) e^{j2\pi f\tau} df$$

Uncorrelated Stationary Process

$$\phi(\tau) = \delta(\tau)$$

$$\Phi(f) = 1$$

Power Spectral Density

Average Power of a Stationary RP

$$\phi(0) = \mathcal{E}\{|X_t|^2\} = \int_{-\infty}^{\infty} \Phi(f) df$$

Symmetry of PSD

Show that: $\Phi^*(f) = \Phi(f)$

Cross Correlation Spectrum

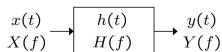
$$\Phi_{XY}(f) = \mathcal{F}\{\phi_{XY}(\tau)\} = \int_{-\infty}^{\infty} \phi_{XY}(\tau) e^{-j2\pi f\tau} d\tau$$

Show that:

$$\Phi_{XY}^*(f) = \Phi_{YX}(f)$$

$$\Phi_{YX}(f) = \Phi_{XY}(-f), \quad \text{For real } X(t) \text{ and } Y(t)$$

Response of LTI Systems



$$y(t) = \int_{-\infty}^{\infty} h(\tau)x(t-\tau)d\tau$$

$$H(f) = \mathcal{F}\{h(t)\}$$

$$\phi_{hh}(\tau) = \int_{-\infty}^{\infty} h(\tau)h(t+\tau)d\tau$$

$$\Phi_{hh}(f) =$$

Mean, ACF and PSD of $Y(t)$

$$m_Y = m_X H(0)$$

$$\begin{aligned}\phi_{YY}(\tau) &= \mathcal{E}\{Y_{t_1} Y_{t_1-\tau}^*\} \\ &= \phi_{hh}(\tau) * \phi_{XX}(\tau)\end{aligned}$$

$$\Phi_{YY}(f) = |H(f)|^2 \Phi_{XX}(f)$$

$$\Phi_{YX}(f) = H(f) \Phi_{XX}(f)$$

Show that: Power spectral densities are Non-negative

Cyclostationary Stochastic Processes

Definition

Cyclo-stationary processes are non-stationary processes whose statistical averages are periodic.

Assume a digital communication signal expressed as

$$X(t) = \sum_{n=-\infty}^{\infty} a[n]g(t - nT)$$

Mean of $X(t)$: **Show that the mean is periodic**

$$m_X(t) = \sum_n \mathcal{E}\{a[n]\}g(t - nT) = m_a \sum_n g(t - nT)$$

ACF of $X(t)$: **Show that the ACF is periodic**

$$\phi_{XX}(t + \tau, t) = \mathcal{E}\{X(t + \tau)X^*(t)\} = \dots$$

Bandpass and Lowpass Random Processes

Refer to [Proakis, Section 2.9]



J. Proakis

Digital Communications, 5th Edition.

McGraw Hill.

Thank You

Questions?

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