

Discrete Probability Distribution Review

- Bernoulli Trials
- Binomial Distribution
- Geometric Distribution
- Poisson Distribution
- Poisson Process



Siméon Denis Poisson
(1781-1840)_{France}

“Life is good for only two things:
discovering mathematics and
teaching mathematics.”

Applications of Poisson_{Review}

- ▶ **Context:** number of events occurring in a fixed period of time
 - ▶ Events occur with a known average rate and are **independent**
- ▶ Poisson distribution is characterized by the **average rate** λ
 - ▶ The average number of arrival in the fixed time period.
- ▶ **Examples**
 - ▶ The number of cars passing a fixed point in a 5 minute interval. **Average rate:** $\lambda = 3$ cars/5 minutes
 - ▶ The number of calls received by a switchboard during a given period of time. **Average rate:** $\lambda = 3$ call/minutes
 - ▶ The number of message coming to a router per second
 - ▶ The number of travelers arriving to the airport for flight registration

Poisson Distribution_{Review}

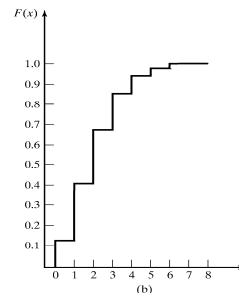
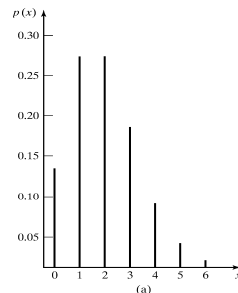
- The Poisson distribution with the average rate parameter λ

$$\text{PMF: } p(k) = P(X = k) = \begin{cases} \frac{\lambda^k}{k!} \exp(-\lambda) & \text{for } k = 0, 1, 2, \dots \\ 0, & \text{otherwise} \end{cases}$$

$$\text{CDF: } F(k) = P(X \leq k) = \sum_{i=0}^k \frac{\lambda^i}{i!} \cdot \exp(-\lambda)$$

Expected value: $E[X] = \lambda$

Variance: $V[X] = \lambda$



Poisson R.V. Review

- Ex. A production line with .4 percent of its items are defective , n=500 items are taken for a quality control. What is the probability that 0, 1, 3 items of them are defective
- That is $X \sim b(500, 0.004)$ approx. To Poisson

$$P(X = k) = \frac{(\lambda)^k}{k!} e^{-\lambda}$$

$$\lambda = 500 * .004 = 2$$

$$P(x = 0) = e^{-2}$$

$$P(x = 1) = 2e^{-2}$$

$$P(x = 3) = \frac{4}{3} e^{-2}$$

Example: Poisson Review

- The number of cars that enter the parking follows a Poisson distribution with a mean rate equal to $\lambda = 20$ **cars/hour**
 - The probability of having **exactly 15 cars** entering the parking in **one hour**:

$$p(15) = P(X = 15) = \frac{20^{15}}{15!} \cdot \exp(-20) = 0.051649$$

Poisson Distribution [Discrete]

- Example: A computer repair person is “beeped” each time there is a call for service. The number of beeps per hour \sim Poisson($\alpha = 2$ per hour).

- The probability of three beeps in the next hour:

$$P(X = 3) = \frac{(\lambda)^k}{k!} e^{-\lambda} = \frac{(2)^3}{3!} e^{-2} = 0.18$$

Or

$$P(X = 3) = F(3) - F(2) = 0.857 - 0.677 = 0.18$$

- The probability of two or more beeps in a 1-hour period:

$$P(X \geq 2) = 1 - p(0) - p(1) = 1 - F(1) = 0.594$$

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Poisson

- Mass function
$$p(x) = \begin{cases} \frac{e^{-\lambda} \lambda^x}{x!} & x \in \{0, 1, \dots, t\} \\ 0 & \text{otherwise} \end{cases}$$
- Algorithm

1. Let $a = e^{-\lambda}$, $b = 1$, $i = 0$

2. Generate $U_{i+1} \sim U(0,1)$ and replace b by bU_{i+1} .

If $b < a$, return $X = i$. Otherwise go to step 3.

3. Let $i = i + 1$ and go back to step 1

- Rather slow. No very good algorithm for Poisson distribution

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Poisson Process

- A stochastic process $\{N(t), t \geq 0\}$ that counts the number of events until time t
- **Examples:**
 - The arrival of "customers" is commonly modeled as a Poisson process in the study of simple queuing systems.
 - The number of web pages requests arriving at a server (except for unusual circumstances such as coordinated denial of service attacks.)

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Stationary & Independent Increments

- A counting process has ***independent increments*** if, for any $0 \leq s < t \leq u < v$, $N(t) - N(s)$ is independent of $N(v) - N(u)$
That is, the numbers of events that occur in nonoverlapping intervals are independent random variables.
- A counting process has ***stationary increments*** if the distribution of $N(t) - N(s)$ depends only on the length of the time interval, $t - s$

Stationary Poisson Process

- **Also known as homogeneous Poisson process** is characterized by a **CONSTANT rate parameter λ** , also known as **intensity**, such that the number of events in time interval $(t, t + \tau]$ follows a Poisson distribution with associated parameter $\lambda \cdot \tau$.

- Formally, A counting process $\{N(t), t \geq 0\}$ is a (Stationary) Poisson process with mean rate λ if:

for $t \geq 0$ and $n = 0, 1, 2, \dots$

$$\text{PMF: } p[N(t + \tau) - N(t) = n] = p[N(\tau) = n] = \frac{(\lambda \cdot \tau)^n}{n!} \exp(-\lambda \cdot \tau)$$

- $N(t + \tau) - N(t)$ describes the number of events in time interval $(t, t + \tau]$
- The mean and the variance are equal $E[N(\tau)] = V[N(\tau)] = \lambda \cdot \tau$

Stationary Poisson Process

- **Properties of Poisson process**

- Arrivals occur one at a time (not simultaneous)
- $\{N(\tau), \tau \geq 0\}$ has **stationary increments**, which means $N(t) - N(s) = N(t - s)$

The number of arrivals in time s to t is also Poisson-distributed with mean $\lambda \cdot (t - s)$

- $\{N(\tau), \tau \geq 0\}$ has **independent increments**

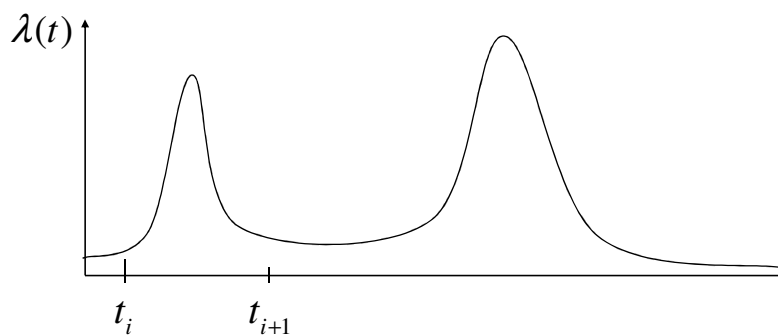
Generating a Poisson Process

- Stationary with rate $\lambda > 0$
- Time between events $A = t_i - t_{i-1}$ are IID exponential
- Algorithm
 1. Generate $U \sim U(0,1)$
 2. Return $t_i = t_{i-1} - \left(\frac{1}{\lambda}\right) \ln U$

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Nonstationary Case

- Can we simply generalize?



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Thinning Algorithm

1. Set $t=t_{i-1}$
2. Generate U_1, U_2 I.I.D. $U(0,1)$
3. Replace t by
$$t - \left(\frac{1}{\lambda^*} \right) \ln U_1, \text{ where } \lambda^* = \max_t \{ \lambda(t) \}$$
4. If $U_2 \leq \lambda(t) / \lambda^*$ return $t_i = t$. Otherwise, go back to step 2.